

FA leads to a decomposition of a space–time signal in the following form:

$$X(M, t) = \sum_{j=1}^p A_j(M)B_j(t) \cos(\omega_j t + k_{j,x}X + k_{j,y}Y),$$

where

$X(M, t)$ is the space–time signal,

M is the space variable corresponding to the zonal (X) and meridional (Y) directions,

t is the time variable,

$A_j(M)$ is the module of the CEOF j (CEOF n° j),

$B_j(t)$ is the module of the complex principal component j (CPC n° j),

$\omega_j = 2\pi/T_j$ is the pulsation of the CPC n° j (corresponding to the period T_j),

$\omega_j t = \varphi_j(t)$ represents the time phase of the CPC n° j ,

$k_{j,x} = 2\pi/\lambda_{j,x}$ corresponds to the zonal wavenumber of the CEOF n° j ,

$k_{j,y} = 2\pi/\lambda_{j,y}$ corresponds to the meridional wavenumber of the CEOF n° j , and

$k_{j,x}X + k_{j,y}Y = \Phi_j(M)$ represents the space phase of the CEOF n° j .

Looking at a traveling wave $X(M, t) = A(M)B(t) \cos[\varphi(t) + \Phi(M)]$ as a matter of evidence, the previous decomposition is particularly relevant for this kind of phenomenon.

To compute CEOFs, one can use two different methods. The first one uses the Hilbert's transform (Barnett 1983). This leads one to use a complex form of the real signal and consequently to compute the complex covariance matrix of the complex signal. CEOFs are given by eigenvectors (and associated eigenvalues) of the complex covariance matrix.

The second way, corresponding to the method proposed by Wallace and Dickinson (1972), uses time cross-spectrum computations. Then, integrating the cross-spectrum matrix over frequencies, we can obtain the complex covariance matrix. The different methods have been compared in Déqué (1986), and we chose to compute CEOFs using the cross-spectrum matrix. More precisely, we calculated the cross-spectrum using a sample method taking into account some advantages highlighted by Déqué (1986). So, we have split the six months of observations per year in eight samples corresponding to 92 observations each, that is to say, 23 days by sample (four observations per day). The length of each sample has been chosen considering the characteristic periods of AEWs (around 4 days) and the problems linked to the Fourier decomposition (namely, we must have enough periods in each sample in order to have a rather good estimation of the waves).

As recommended by Déqué (1986), we weighted data at each grid point taking into account the spherical surface of the domain. This led us to introduce a weight

that is proportional to the root of the cosine of the latitude.

Just before applying the CEOFA method, we used a classic empirical orthogonal function analysis (EOFA), in order to provide an easier calculation of CEOFs and to look at the ability of this factorial method to identify AEWs. Additionally, we made some sampling sensitivity tests both in space and time domains. Looking at the sensitivity of the methods to the space domain, we tested four different domains. One can see in Figs. 4a and 4b that the patterns of EOFs are quite stable from the larger domain (used for STSA) to the smaller one (finally retained for EOFA and CEOFA). In fact, the main differences between different analyses done are in the rank and percentage of variance corresponding to associated eigenvalues.

In the same way, results from raw and filtered data are also quite comparable both in space (for EOFs and CEOFs, not shown here) and time domains [for principal components (PCs) and CPCs]. As shown in Figs. 5a and 5b, the main difference came from the smoothing effect of the filter in the time domain. It has been noticed that the space phases (not shown here) are less noisy for filtered data.

3. Results from the space–time spectral analysis

a. Wave spectra

Figure 6 shows the space–time spectra (the all-space domain and May–October 1985) of the traveling waves for ECMWF (top panel) and the GCM (bottom panel). These diagrams give the power density versus the frequency (negative for eastward propagating phenomena) in the abscissa and the zonal wavenumber in the ordinate. On both spectra, there are several maxima of power density between 36 and 60 in time (3–5 days) for westward propagating phenomena and 2 and 4 in space (2200–4400 km). This spectral window is consistent with what has been found by other authors (e.g., Reed et al. 1988b) for the AEWs. However, the model variance maxima tend to be located toward lower frequency compared to the analysis variance maxima. Indeed, the secondary maximum of power density around 37 (5 days) is larger in the spectrum of the simulated data (Fig. 6). The order of magnitude of the model variance cumulated in the AEW spectral window is the same as that of the analysis variance (this topic will be discussed in more detail in section 4). Interestingly, there are also local maxima on both spectra for much lower frequencies of around ± 5 in time (37 days, for eastward and westward propagating oscillations) and for the first wavenumber. The same spectra computed for standing plus traveling waves (not shown) allow us to say that traveling wave variance represents 23% of the variance of all the waves, but more than 70% of the AEW variance. We have noticed the same kind of pattern for the other years, indicating a good similarity between anal-

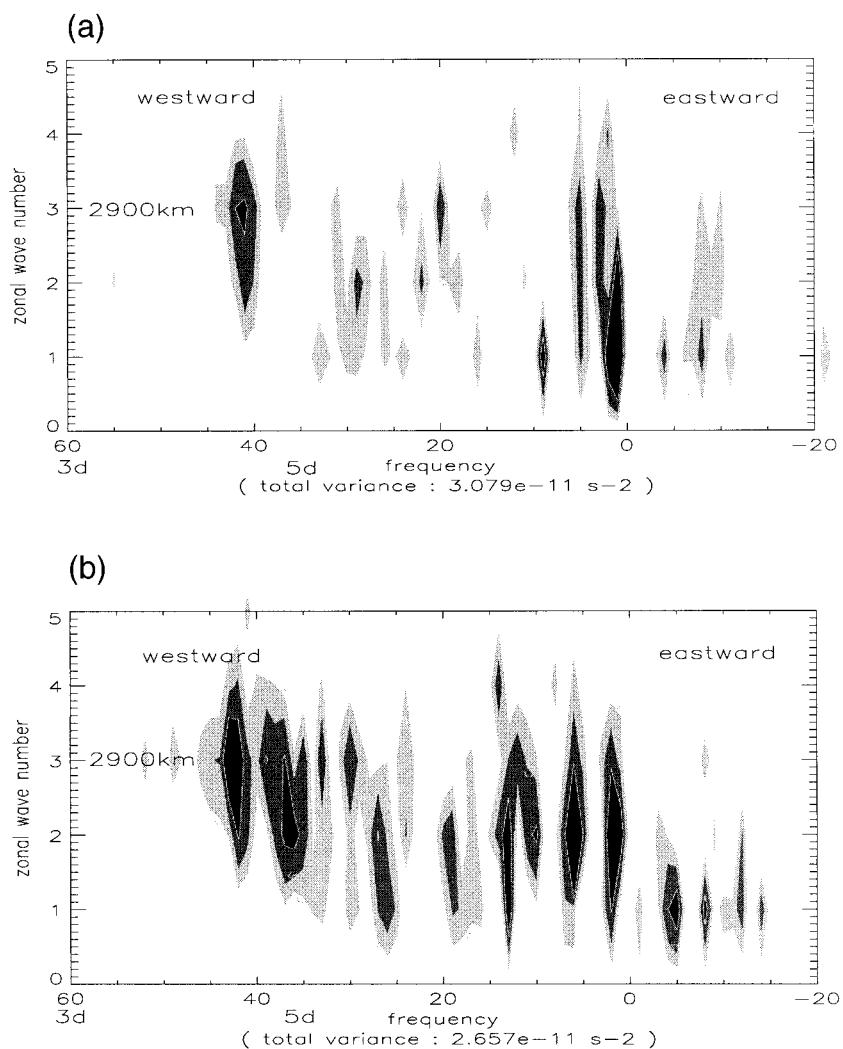


FIG. 6. Space-time spectra for the relative vorticity at 850 hPa (traveling waves only) from May to Oct 1985, expressed in thousandth of the total variance (isolines: 2, 3, 8); (a) ECMWF, (b) CNRM GCM. Abscissa are given in time harmonics (maximum period of 184 days, negative for eastward propagation), ordinate in zonal harmonics (maximum wavelength of 8800 km).

AEJ may explain the fact that the simulation gives less variance south of 12°N compared to the analysis.

On the other hand, another major discrepancy between the simulation and the analysis is the fact that there is more variance east of 10°E in the model, as also indicated by Druyan and Hall (1994). At this stage, it is difficult to explain this discrepancy. There is a lack of observations in this region, therefore the analysis might not be reliable.

c. Energetics

Zonal wind–meridional wind (u, v) and meridional wind–temperature (v, T) covariances at 850 hPa have been computed in the AEW space–time spectral window, with the help of a space–time cross-spectrum calculation (Hayashi 1982). Figure 9 shows these covariance contours at

850 hPa, in July 1985 [(u, v) for ECMWF and the GCM, and (v, T) for the GCM]. It would have been more relevant to consider the 700-hPa level, the closest to the AEJ level, to compute the (u, v) covariance, but the simulated data were not available at this level. Nevertheless, we notice a close correspondence between the simulated and analyzed (u, v) covariance. A zone of positive (u, v) values lies south of the AEJ (zone in which $\partial u/\partial y$ is negative) in the western part of Africa, meaning a contribution to a barotropic transfer of kinetic energy from the jet to the waves. Furthermore, there is a zone of negative (v, T) values lying around the location of the AEJ (zone in which $\partial T/\partial y$ is positive), meaning a contribution to a baroclinic transfer of potential energy from the jet to the waves over that region (between 12° and 18°N, rather to east of the region associated with the barotropic transfer of kinetic energy). This pattern is very close to the one obtained by Reed et

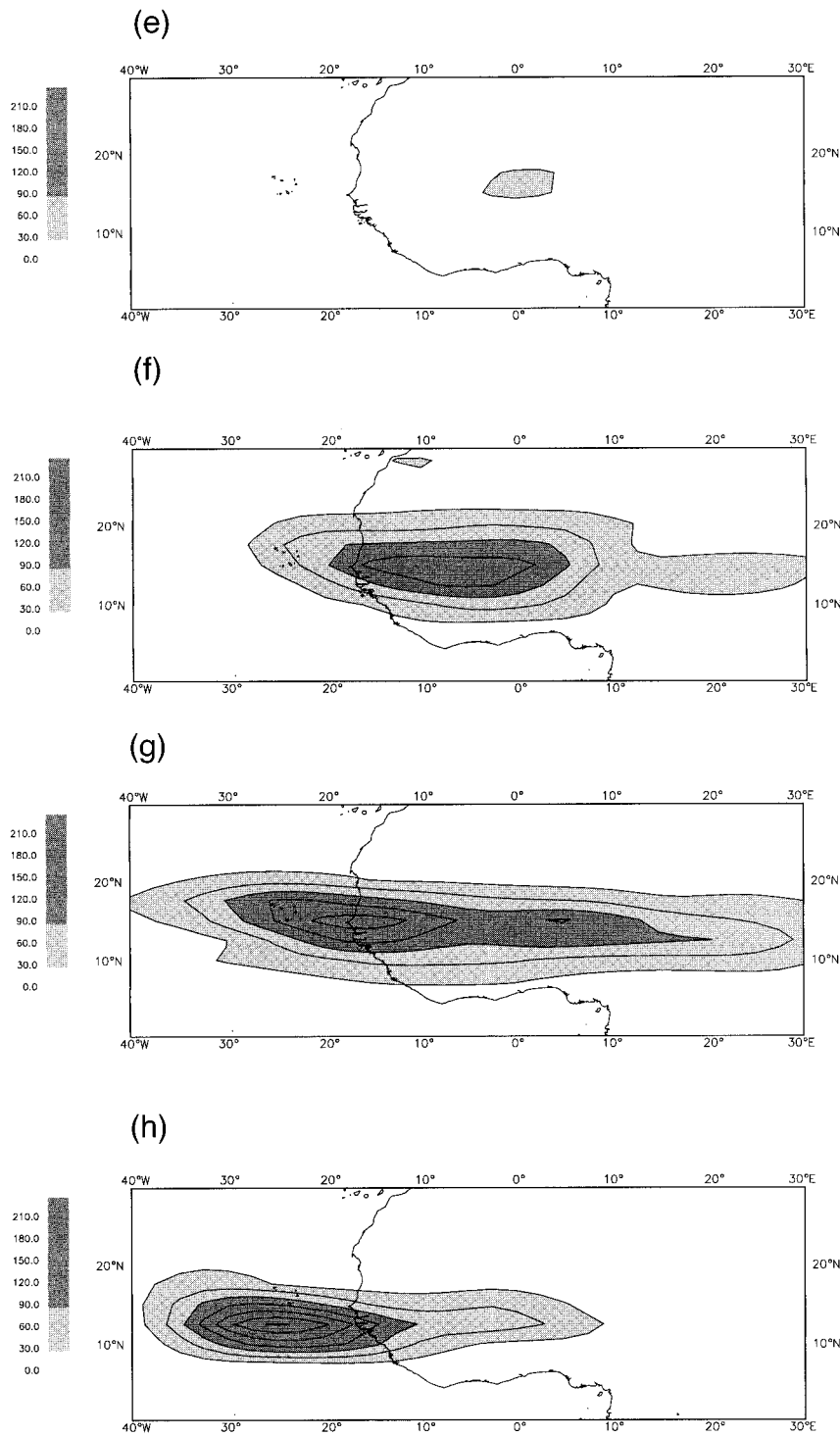


FIG. 8. (Continued)

over West Africa (see next section) strengthens the following interpretation: the northern single-track mode (i.e., mode 1) should be preferentially representative for AEWs associated with dynamical effects (i.e., with a predominant northerly component), while the dual-track

mode (i.e., mode 2) should catch the majority (in comparison with mode 1) of the AEWs associated with diabatic effects. In other words, the dual-track mode composes the majority of the AEWs characterized by a significant southerly component but also some AEWs char-

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